

A 20-Year Record of the Western Prairie Fringed Orchid (*Platanthera praeclara*): Population Dynamics and Modeling of Precipitation Effects

Authors: Morrison, Lloyd W., Haack-Gaynor, Jennifer L., Young, Craig C., and DeBacker, Michael D.

Source: Natural Areas Journal, 35(2) : 246-255

Published By: Natural Areas Association

URL: <https://doi.org/10.3375/043.035.0205>

The BioOne Digital Library (<https://bioone.org/>) provides worldwide distribution for more than 580 journals and eBooks from BioOne's community of over 150 nonprofit societies, research institutions, and university presses in the biological, ecological, and environmental sciences. The BioOne Digital Library encompasses the flagship aggregation BioOne Complete (<https://bioone.org/subscribe>), the BioOne Complete Archive (<https://bioone.org/archive>), and the BioOne eBooks program offerings ESA eBook Collection (<https://bioone.org/esa-ebooks>) and CSIRO Publishing BioSelect Collection (<https://bioone.org/csiro-ebooks>).

Your use of this PDF, the BioOne Digital Library, and all posted and associated content indicates your acceptance of BioOne's Terms of Use, available at www.bioone.org/terms-of-use.

Usage of BioOne Digital Library content is strictly limited to personal, educational, and non-commercial use. Commercial inquiries or rights and permissions requests should be directed to the individual publisher as copyright holder.

BioOne is an innovative nonprofit that sees sustainable scholarly publishing as an inherently collaborative enterprise connecting authors, nonprofit publishers, academic institutions, research libraries, and research funders in the common goal of maximizing access to critical research.

A 20-Year Record of the Western Prairie Fringed Orchid (*Platanthera praeclara*): Population Dynamics and Modeling of Precipitation Effects

Lloyd W. Morrison^{1,2,3}

¹National Park Service
Heartland Inventory and
Monitoring Program
6424 W. Farm Road 182
Republic, MO 65738

Jennifer L. Haack-Gaynor¹

Craig C. Young¹

Michael D. DeBacker¹

²Department of Biology
Missouri State University
901 S. National Avenue
Springfield, MO 65897

³ Corresponding author:
LloydMorrison@MissouriState.edu; 417-
836-3119

ABSTRACT: Western prairie fringed orchid (*Platanthera praeclara*) populations may be highly variable over time. It is thought that soil moisture, and potentially disturbance, play important roles in determining abundance and flowering in this species. We present data for a *P. praeclara* population in southwestern Minnesota that has been monitored annually for 20 years. The number of flowering orchids varied over this period from 0 to 722. We provide an empirical test of an earlier model of precipitation effects during phenological life stages of the orchid, finding that the model was no better than a simple null model that ignored variability in precipitation. We re-examined the relationship between number of flowering orchids and precipitation during phenological life stages based on a modern information theoretic (AIC), multimodel inference approach, and a larger data set. The models indicate the importance of precipitation during three phases of orchid life history: (1) mature growth in the previous year; (2) postsenescence; and (3) emergence, explaining >70% of the variation in the number of flowering orchids. We also evaluated the effect of prescribed burns on this orchid population. Although we found no effect of fire on the number of flowering orchids, plants were shorter in burn years. This difference was not significant, however, once variation in precipitation was taken into account.

Index terms: AIC, model selection, *Platanthera praeclara*, prescribed burn, western prairie fringed orchid

INTRODUCTION

The western prairie fringed orchid (*Platanthera praeclara* Sheviak and Bowles) is a perennial, herbaceous orchid native to tallgrass prairies from south-central Canada through the western central lowlands and eastern Great Plains of the United States (US Fish and Wildlife Service 1996). This species was listed as threatened in 1989 (US Fish and Wildlife Service 1989). Approximately 90% of all extant plants in North America occur in North Dakota, Minnesota, and Manitoba, Canada (Sheviak and Bowles 1986; US Fish and Wildlife Service 1996). Population sizes range from tens to thousands of plants; however, most remaining populations consist of <50 individuals (US Fish and Wildlife Service 1996). The orchid's decline is linked with prairie conversion and alteration of hydrological regimes associated with agricultural modification (US Fish and Wildlife Service 1996).

Platanthera praeclara typically emerges in mid-April to late May. Two aboveground growth forms exist: (1) vegetative, consisting of only a few leaves, and (2) flowering, with inflorescences consisting of a spike of up to 40 creamy white flowers (Bowles 1983). Plants originate either from perennating buds that developed on tubers during the previous year, or from germinated seeds. Germination and development of protocorms require a symbiotic association with mycorrhizal

fungi, particularly *Ceratorrhiza* (Bowles 1983; Sharma et al. 2003). Vegetative reproduction occurs infrequently following the production of multiple primary tubers and the death of the primary tuber from the previous year. Both perennating buds and flowering buds form during the previous year. Flowering typically occurs from mid-June to late July. Individual flowers persist for up to ten days, and a single inflorescence produces flowers for up to three weeks (Bowles 1983). A relatively large proportion of plants may be dormant or vegetative in any given year.

Platanthera praeclara populations, at least the flowering or nondormant proportion, may be highly variable over time (Sieg and King 1995; Sather and Anderson 2010). *Platanthera praeclara* is usually found in mesic swales or draws, and soil moisture appears to affect *P. praeclara* dormancy and flowering (Bowles 1983; Bowles et al. 1992; Bray and Wilson 1992; Sieg and King 1995). Moisture may provide an important cue either in the current or previous growing season, because flower bud initiation occurs during the previous growing season (Bowles 1983). Willson et al. (2006) used stepwise multiple linear regression to select a model predicting the number of flowering orchids as a function of cumulative precipitation during six phenological stages of the orchid spanning two growing seasons. Although the model fit the data used to construct the model relatively well, the predictive ability of the model has not been formally evaluated.

In addition to moisture, disturbances that remove competing species may also be important in determining the number of flowering orchids (Bowles 1983). Some mostly anecdotal observations (Sheviak 1974; Currier 1984) and a limited experimental study (Pleasants 1994) suggest fire may increase *P. praeclara* flowering.

Here we present data spanning 20 years (1993–2012) for a *P. praeclara* population in southwestern Minnesota. We formally test the model of Willson et al. (2006), which was constructed from a 10-year dataset (1995–2004) from this same population, and apply a more modern information theoretic approach to evaluate the effect of precipitation on *P. praeclara*. We also evaluate the effect of a limited number of prescribed burns on this population.

We pose the following questions: (1) How variable was the *P. praeclara* population over the 20-year period? (2) How does precipitation affect the number of flowering orchids? and (3) Is there any evidence of a fire effect, once the effect of precipitation is taken into account?

METHODS

Study Area

A *P. praeclara* population exists at Pipestone National Monument in southwestern Minnesota (44° 0.799 N, 96° 19.518 W). The population occurs in mesic prairie habitat that was part of the original 46 hectares designated as a national monument in 1937, which had

escaped cultivation. The soil type consists of Ihlen and Ihlen-Rock outcrop complex (Becker et al. 1986). Ihlen soils are well drained and consist of 20–40 inches of loess over Sioux quartzite on relatively level surfaces. The Ihlen-Rock complex includes rock outcrops within thinner Ihlen soils. Shallow drainages are found in both soil types. Wildfires occurred until the early 1940s, when relatively effective fire suppression began. Resource managers initiated prescribed burning in the orchid habitat in 1973.

Sampling Methodology

Surveys were conducted for *P. praeclara* annually from 1993 to 2012. Monitoring took place during peak flowering, which always occurred in July, with the exception of 2012, in which peak flowering occurred during the last week of June. In 1993 and 1994, a haphazard search of the area was conducted by two observers. Beginning in 1995, a systematic search of the area was conducted by a team of four to eight observers walking in tandem 2 m apart (Willson 2000; Young et al. 2007). Locations of flowering orchids were recorded with a GPS unit. Accuracy of GPS positions was within 5 m through 2006, and within 2 m for 2007 and later years. Plant height and the number of flowers and buds were also recorded. Given the more thorough nature of the surveys beginning in 1995, the number of flowering orchids for 1993 and 1994 could be underestimates, and plant height and number of flowers/buds could also be biased. Thus, we include these data for descriptive purposes, but for modeling effects of precipitation we

include only years in which the same monitoring methodology was used (as did Willson et al. 2006). Prescribed burns were conducted in four years of this study: 1994, 1997, 2002, and 2009, always in May. The entire orchid habitat was burned as uniformly as possible.

Statistical Analyses

We evaluated the effect of precipitation on number of flowering plants during six phenological stages as identified by Wolken (1995) (Table 1). We first tested the model obtained by Willson et al. (2006) (referred to as the Willson et al. model in the results): (Number of flowering plants = 196.73 + 7.28 precipitation during senescence/bud development – 9.30 precipitation during dormancy), which explained 77% of the variability in flowering plant number for the years used to construct the model. We compared the observed number of flowering plants to the number predicted by the model in each of the eight years following publication of the model, and calculated the average absolute difference between observed numbers and the model predictions. We also determined for each year, as a null model of no precipitation effect, the difference between the model prediction and the average number of flowering orchids over the same eight years on which the model was based. We calculated the average absolute difference between observed numbers and this long-term average.

Next, we re-evaluated the relationship between number of flowering plants and precipitation in the specified phenological

Table 1. Six phenological stages of *P. praeclara* as identified by Wolken (1995). Time period refers to the dates over which precipitation was accumulated for each stage.

<i>P. praeclara</i> phenological stage	Time period
Mature growth in year before flowering	June 1 – July 31 of previous year
Senescence/bud development	August 1 – 31 of previous year
Postsenescence	September 1 – 30 of previous year
Dormancy	October 1 – December 31 of previous year and January 1 – March 31 of current year
Emergence	April 1 – May 31 of current year
Mature growth in year of flowering	June 1 – July 31 of current year

stages by employing an information-theoretic approach. In this method, inference is based on Kullback-Liebler (K-L) information (Kullback and Liebler 1951). Akaike's Information Criterion (AIC) (Akaike 1973) was used as an estimator of relative K-L information (Burnham and Anderson 2002). We conducted all subsets regression, and used AIC to evaluate the regression models. Because the sample size was small, we calculated AIC adjusted for small sample size (AIC_c) as:

$$AIC_c = AIC + 2K(K + 1)/(n - K - 1)$$

where K is the total number of parameters in the model (including the intercept and residual variance) and n is the sample size. Because AIC_c is on a relative scale, differences in AIC_c values were calculated as:

$$\Delta AIC_{ci} = AIC_{ci} - \text{minimum } AIC_c$$

for all candidate models, where i indicates the i^{th} model. The best approximating model was the one with the lowest AIC_c score. Models with $\Delta AIC_c \leq 2$ have "substantial" support, however, and would serve nearly equally well in approximating the information in the data set (Burnham and Anderson 2002). We report all models with $\Delta AIC_c \leq 2$.

A total of 18 years of data obtained with a consistent methodology was available (1995–2012), although prescribed fires occurred in three years (1997, 2002, and 2009) (Appendix). Because fires may influence flowering of the orchid, an analysis was conducted excluding the burn years ($n = 15$ years), and a separate analysis was conducted including all years of data ($n = 18$).

We conducted parallel analyses using plant height and number of flowers/buds as response variables in place of number of flowering plants, to evaluate whether there was any evidence that precipitation in the six phenological stages affected these variables. As in the analyses focusing on number of flowering orchids, analyses for plant height and number of flowers/buds were conducted for all years of data (1995 to 2012) and excluding burn years. Because no orchids were present in 1998, this year

was excluded from all analyses of plant height and number of flowers/buds. Data from 1993 and 1994 were not included; because of the haphazard nature of the sampling in those years, shorter orchids, or those with fewer flowers, may have been missed.

Durbin-Watson tests were used to evaluate the linear models with substantial support, to determine whether the error terms were serially correlated. Positively autocorrelated error terms often result from a model that is missing key variables (Kutner et al. 2005).

We evaluated whether the prescribed burns affected number of flowering orchids, plant height, or number of flowers/buds with t -tests and one-way analysis of covariance (ANCOVA). We evaluated potential effects of burns both in the year of the burn and the following year, in separate analyses, since flower bud initiation occurs in the year before flowering (Bowles 1983), and removal of litter from fire may affect orchids in the year after the burn (Pleasants 1994).

We obtained daily precipitation records for the Pipestone weather station (216565) from the National Oceanic and Atmospheric Administration (NOAA) Global Historical Climatology Network (GHCH) (<http://gis.ncdc.noaa.gov/map/viewer/>). This weather station was located ~600 m from the orchid population. A small fraction of the precipitation data was missing—0.6% of the total records—and in these cases, the average of three surrounding weather stations was substituted. The three weather stations used were Tyler (218429; ~20 miles to the northeast), Flandreau (392984; ~13 miles to the west), and Lake Wilson (214534; ~18 miles to the east). Fifty-five percent of the missing values were zeros.

All subsets regression and AIC evaluation were conducted using SAS® software (SAS Institute Inc. 2008). Inspection of diagnostic plots revealed no major departures from regression assumptions such as normality or constancy of error variance. Durbin-Watson tests, t -tests, and AVOVAs were done with IBM SPSS Statistics software (IBM Corp. 2011). The sequential

Bonferroni method (Rice 1989) was used to correct for multiple comparisons when necessary.

RESULTS

Demographics

The spatial distribution of the orchid was generally clustered around two low, wet areas (Figure 1). The number of flowering orchids was variable, but generally <150 were observed in the first 18 years of the study, with >200 individuals found in only one of those years (2003) (Figure 2A). Numbers were often quite low, with no orchids being found in 1998, and ≤ 3 in 1997 and 2007. In contrast, the number of flowering orchids in both of the final two years of the study was >600. The mean number of flowers/buds varied between 5 and 10 across the entire data record, but revealed a relatively steady decline over the first 17 years of the study, before rebounding in the last three years (Figure 2B). Mean plant height was greatest in the first four years of the study, and highly variable (albeit lower) afterwards (Figure 2C).

The number of flowering orchids was not significantly correlated with either mean plant height or mean number of flowers/buds. Mean plant height and mean number of flowers/buds, however, were positively correlated ($r = 0.49$; $P = 0.032$), although this was not significant if corrected for multiple comparisons by the sequential Bonferroni method ($n = 3$ comparisons).

Effects of Precipitation

The Willson et al. model yielded predictions that were usually not close to the observed numbers of flowering plants (Table 2). Interestingly, burn years were excluded from the data used to construct the Willson et al. model, yet the most accurate prediction occurred in 2009, the only burn year among these comparisons. The average number of flowering plants in the data set used in the Willson et al. model was 87. If one were to use this simple average as a predictor for the number of

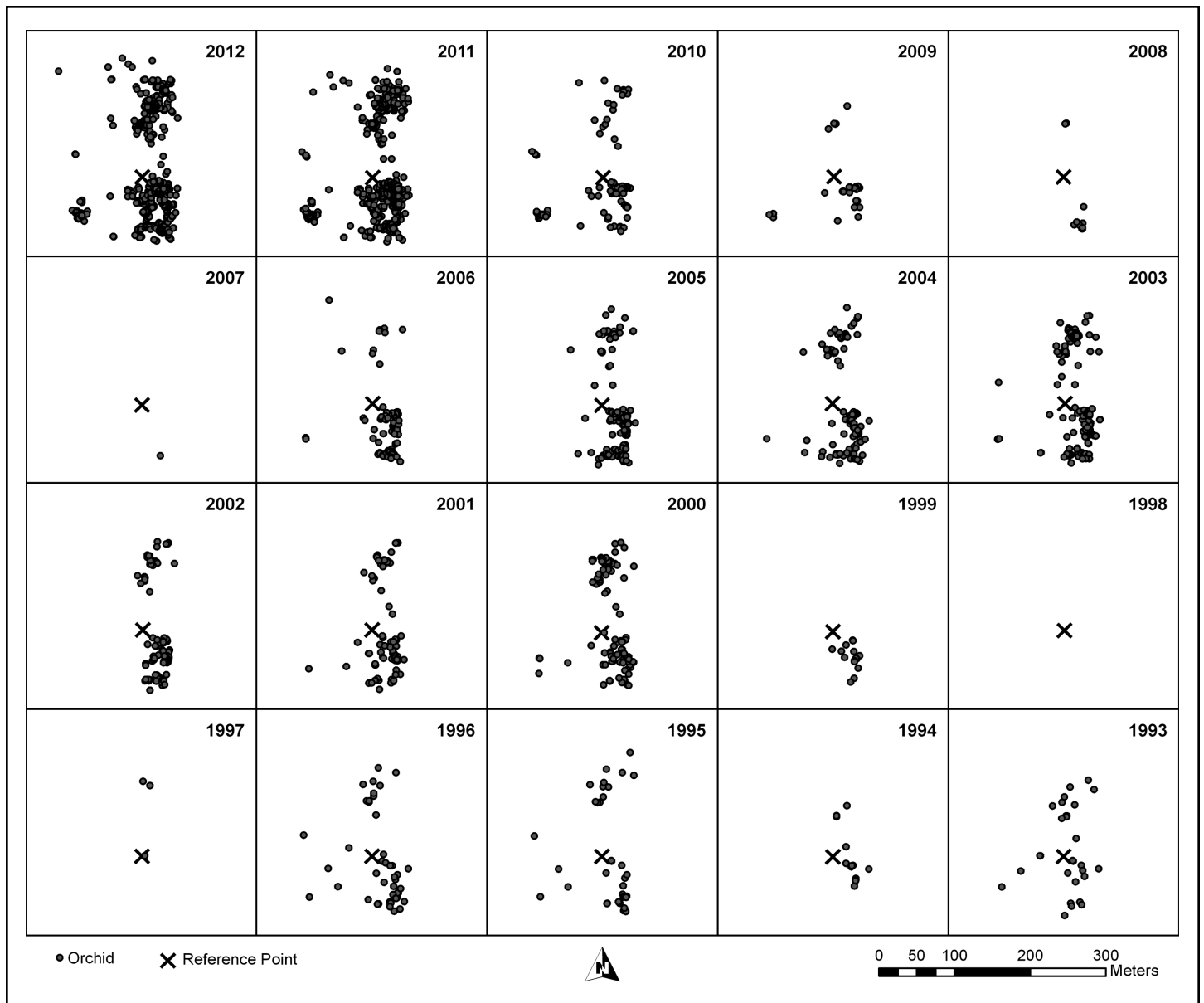


Figure 1. Map panels showing distribution of flowering orchids in each of 20 years (1993–2012) at Pipestone National Monument, Minnesota. Accuracy of GPS positions was within 5 m through 2006, and within 2 m for 2007 and later years.

flowering plants in years succeeding the model (2005–2012), the average absolute difference between the observed number and the average would be 187. The average absolute difference between the observed number and the number predicted by the Willson et al. model for this time period is 194. Thus, over this time period, simply using the historical average would be a slightly better predictor of orchid numbers than use of the model.

When the larger data set was analyzed by

an information theoretic approach, and burn years excluded ($n = 15$ years), the best approximating model based on AIC_c scores included two categories: (1) mature growth in the previous year (the year before flowering); and (2) emergence (Table 3). Two other models, however, were essentially as good. These three models always contained some subset of three categories: (1) mature growth in the previous year; (2) postsenescence; and (3) emergence (or all three). The model containing the predictor variables senescence/bud development and

dormancy (the two variables in the Willson et al. (2006) model) had $\Delta AIC_c = 16.42$. A model with $\Delta AIC_c > 10$ has “essentially no empirical support” (Burnham and Anderson 2002).

If burn years are included ($n = 18$ years), the same best approximating model was obtained (with somewhat different coefficients) (Table 3). Only one other model (with all three of the categories) was essentially as good, however. The relatively small differences in the coefficients are

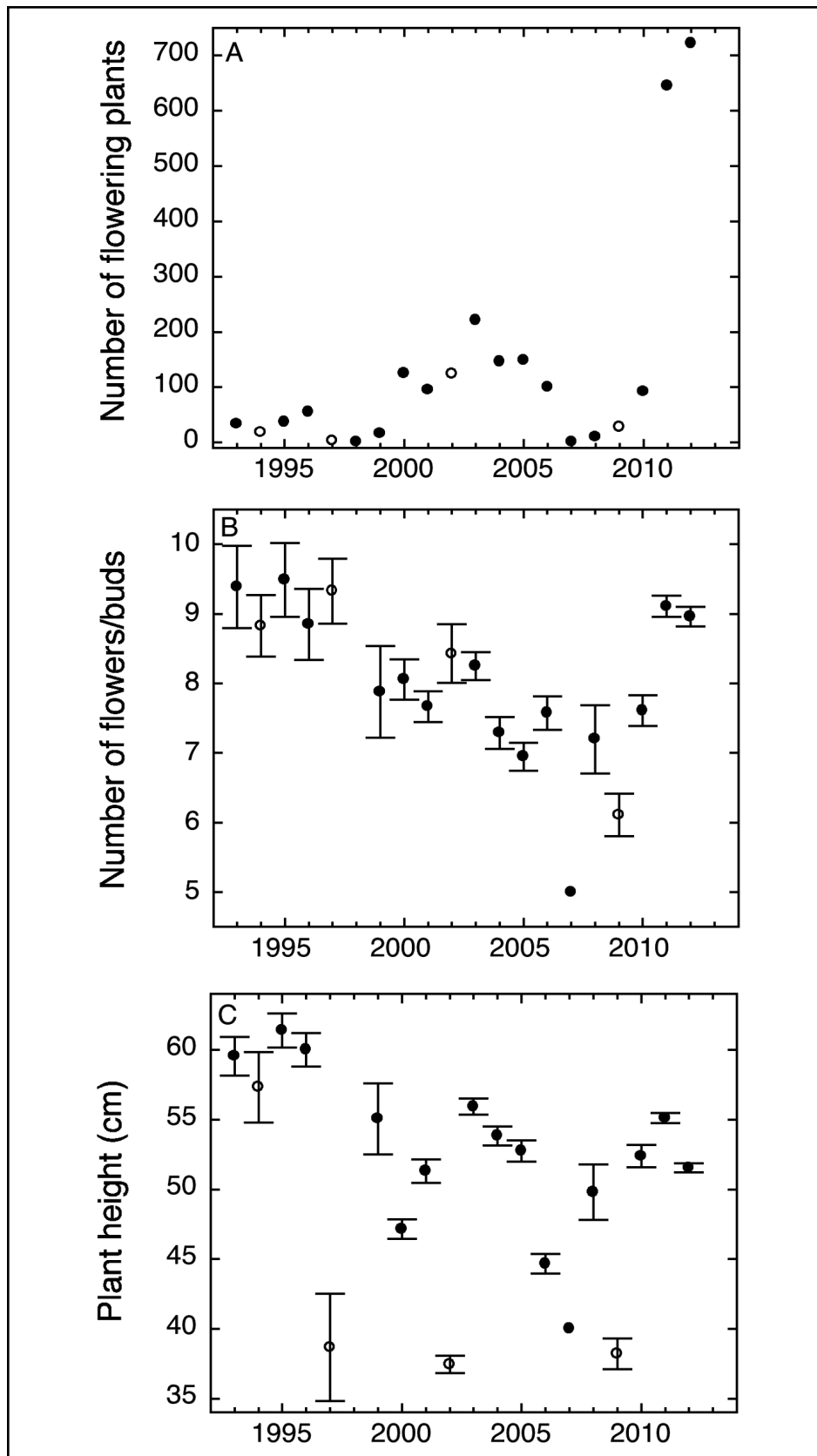


Figure 2. (A) Number of flowering plants from 1993 to 2012 at Pipestone National Monument, Minnesota. (B) Mean number of flowers and buds (Error bars = 1 SE). (C) Mean plant height in cm (Error bars = 1 SE). In all panels, open circles indicate years in which a prescribed burn occurred. Closed circles indicate years without fire.

likely to result more from model selection uncertainty due to small sample sizes rather than an effect of fire. Again, the model containing the predictor variables senescence/bud development and dormancy performed poorly, with $\Delta AIC_c = 19.26$.

All Durbin-Watson statistics were in the indeterminate range, yielding inconclusive test results (D ranged from 0.95 to 1.30). In principle, this is due to lack of enough cases to obtain a definitive result.

Analyses of models with plant height and number of flowers/buds as response variables revealed that relatively little of the observed variability in either response variable was explained by variability in predictor variables in the most parsimonious models. Models with $\Delta AIC_c \leq 2$ yielded coefficients of determination of ≤ 0.38 for number of flowers/buds and ≤ 0.24 for height in years with no burns ($n = 14$); maximum values of R^2 for all years ($n = 17$) were ≤ 0.14 for number of flowers/buds and ≤ 0.42 for height. We do not present the models here since all were relatively poor.

Effects of Fire

There was a marginally significant ($P = 0.088$) effect of fire on the number of flowering orchids, which was lower in burn years than nonburn years (43.3 ± 54.9 vs. 152.9 ± 216.7 , mean \pm SD). Because precipitation was observed to have important effects on the number of flowering orchids in the regression models, precipitation values during the three phenological stages found to be important (mature growth in the previous year, emergence, and postsenescence) were included as covariates in an ANCOVA. When these variables were taken into account, there was no effect of fire ($F = 0.09$, $P = 0.77$, $df = 1$).

The number of flowers/buds was not different in burn years vs. nonburn years ($P \gg 0.05$). Plant height was shorter in years with a prescribed burn (42.9 ± 9.6 vs. 52.7 ± 5.8 (mean \pm SD); $t = -2.63$, $df = 17$, $P = 0.018$). If only years with a consistent, systematic monitoring are analyzed (all

Table 2. Observed number of flowering orchids, number of flowering orchids predicted by the Willson et al. (2006) model, the absolute difference between observed and predicted numbers, and the absolute difference between observed numbers and the average number of flowering orchids (87) over the period covered by the Willson et al. (2006) model, at Pipestone National Monument, Minnesota. A burn occurred in 2009.

Year	Observed number (O)	Predicted number (P)	Absolute difference O – P	Absolute difference O – 87
2005	149	168	19	62
2006	100	24	76	13
2007	1	147	146	86
2008	10	113	103	77
2009	28	23	5	59
2010	92	25	67	5
2011	645	121	524	558
2012	722	112	610	635
Average absolute difference			194	187

DISCUSSION

Effects of Precipitation

Based on an examination of *P. praeclara* root systems, Wolken (1995) proposed a phenological system of above- and below-ground plant development. Our model results indicate the relative importance of precipitation in the various life stages of the previous and current growing season. The importance of precipitation during the previous growing season (June–July) suggests the existence of a lag effect. Some threshold of precipitation, or more likely soil moisture, may be required to promote robust growth of the perennating bud and supporting root system that leads to flowering during the subsequent season. As a more immediate effect, there may exist a precipitation threshold during emergence (April–May) that is required to further support plant development leading to flowering. The periods of mature growth and emergence correspond with periods of greatest photosynthetic activity in *P. praeclara*. This putative mechanism may also explain variation in *P. praeclara* population size (Pleasants 1995) and demographic transitions between flowering, nonflowering, and dormant states (Sieg and King 1995). While our modeling cannot ascertain the relative strengths or interactions of precipitation during these periods, we have detected a signal indicating the relative importance of

years except 1993 and 1994), this difference in plant height is very robust (38.1 ± 0.6 vs. 52.2 ± 5.7 (mean \pm SD); $t = -4.20$, $df = 15$, $P = 0.001$).

The three years with the shortest plants (mean height <40 cm) were all burn years (1997, 2002, and 2009); however, they also happened to be the three years with the least spring precipitation (Figure 3). In fact, there was a significant positive association between plant height and precipitation during the emergence phenological stage

($r = 0.47$, $P = 0.043$). If plant height as a function of prescribed burns is compared using an ANCOVA with precipitation during the emergence stage as a covariate, the effect of fire is no longer significant ($F = 2.97$, $P = 0.104$, $df = 1$).

When similar comparisons were made in the year after fire (i.e., comparing burn years + 1 to all other years), there were no significant differences in either number of flowering orchids, number of flowers/buds, or plant height (all $P > 0.05$).

Table 3. The best approximating model (lowest AIC_c score) and competing models having “substantial” support ($\Delta\text{AIC}_c \leq 2$) for the *P. praeclara* population at Pipestone National Monument, Minnesota. MLL is the maximized log-likelihood. K is the total number of parameters in the model (including the intercept and residual variance). The response variable for all models is number of flowering orchids.

Model/Variables	MLL	AIC _c	ΔAIC_c	K	Adj. R ²	R ²
<i>n</i> = 15 (1995–2012, excluding burn years)						
-343.44 + 13.02 MGprev + 13.06 EM	-70.13	151.91	0	4	0.666	0.714
-436.29 + 8.88 MGprev + 10.12 PS + 17.33 EM	-68.29	152.571	0.67	5	0.715	0.776
-427.181 + 15.89 PS + 22.65 EM	-70.49	152.608	0.7	4	0.65	0.700
<i>n</i> = 18 (1995–2012, including burn years)						
-319.11 + 13.12 MGprev + 12.07 EM	-83.10	177.842	0	4	0.679	0.717
-383.22 + 10.37 MGprev + 8.29 PS + 14.59 EM	-81.41	178.811	0.97	5	0.715	0.765

Note: MGprev = mature growth in the previous year (the year before flowering); PS = postsenescence; EM = emergence

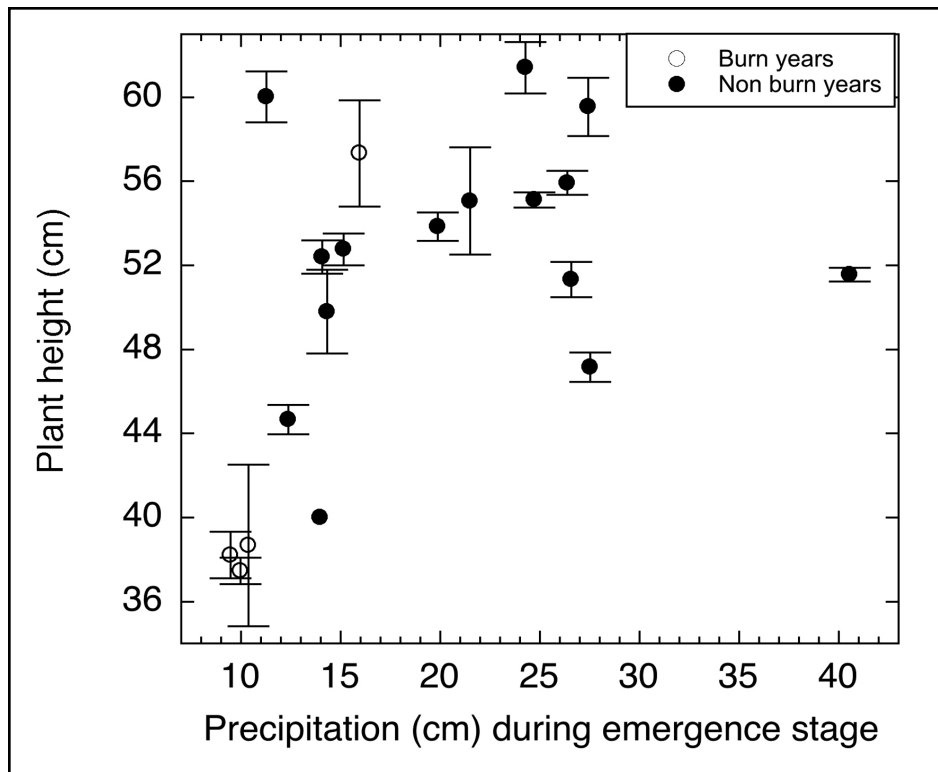


Figure 3. Mean *P. praeclara* height as a function of precipitation during the emergence stage from 1993 to 2012 at Pipestone National Monument, Minnesota. (Error bars = 1 SE).

precipitation that presumably corresponds with distinct phenological stages in *P. praeclara*.

Why did the models obtained using the information theoretic approach differ so greatly from the model obtained by Willson et al. using stepwise multiple linear regression? Numerous issues exist with stepwise model selection procedures, including the problem of multiple hypothesis testing, bias in parameter estimation, inconsistencies among model selection algorithms, and an inappropriate focus on a single best model (Anderson and Burnham 2002; Austin and Tu 2004; Whittingham et al. 2006; Wiegand 2010; Hegyi and Garamszegi 2011).

The most important factor, however, probably has to do with the smaller sample size ($n = 8$) evaluated by Willson et al. Different models may be determined to be “best” for different data sets (from the same population), and this source of model selection uncertainty becomes greater as sample size becomes smaller. The Willson et al. model had essentially no empiri-

cal support compared to the best models determined with larger sample sizes ($n = 15$ and 18). While our sample sizes were larger than those available to Willson et al., they are still relatively small. Instead of constructing a model based on a subset of the data and then testing it on another subset, we used all available data in model selection, in an attempt to obtain the most robust inference possible. Thus, although the models generated may fit these data fairly well, there is still the chance that they are not representative of the overall population.

The available data did not support selection of a single best model. Other methods, such as stepwise regression, are guaranteed to produce a single best model, even though there may exist substantial model selection uncertainty. The inability to identify a single best model is not uncommon when using an information theoretic approach, and reflects the natural complexity of the system, rather than a shortcoming of the method. Robust inferences can still be made from a set of competing models with

substantial support, that is, multimodel inference (Burnham and Anderson 2002). More formal procedures include model averaging, although in this case model averaging would probably not be very useful. The inability to rule out positive autocorrelation in the error terms suggests that remedial measures for autocorrelation (see Kutner et al. 2005) should be explored and transformed models evaluated before attempting to make any inferences from parameter coefficients. Given the caveats above, identifying the important phenological stages, rather than specifying exact parameter coefficients, represents the most important inference from the analyses presented here.

Because of the relatively small sample size, and the fact that we evaluated all possible models, the results of this study should not be considered as confirmatory. Other environmental variables are likely to be important and, like most studies, we consider only simple linear models, although the true relationship may be curvilinear or contain higher-order interactions. The inconclusive results of the Durbin-Watson tests do not preclude the existence of other key variables. Thus, the results of the study should be taken more as an early step toward evaluating the relationship of flowering orchids with environmental variables such as rainfall during different phenological stages, and demonstrating that an earlier published model is not useful for prediction.

A limitation inherent in the information theoretic approach is that this method will only select the best model(s) from a larger set of candidate models. If the candidate models are all poor, then the selected model(s) will also be poor. In the analyses of number of flowering orchids, coefficients of determination (R^2) were always ≥ 0.7 , indicating a relatively large amount of the variability in the number of flowering orchids could be explained by variability in a few precipitation categories. In contrast, in the analyses of plant height and number of flowers/buds, coefficients of determination were always ≤ 0.42 , indicating that the majority of the variability in these two variables was due to other factors outside the scope of these analyses. Finally, our

analyses considered only one *P. praeclara* population. Thus, these results may not be directly generalizable to other populations in other areas.

Effects of Fire

Anecdotal observations in the literature suggest that fire may increase flowering in *P. praeclara* (Sheviak 1974; Bowles 1983; Currier 1984). The effect of fire was difficult to evaluate, since only four prescribed burns occurred over the course of the data record. Moreover, fire in this case is a complex variable, as both the timing and intensity of the burns varied. No consistent increases or decreases were observed in orchid numbers following burns. Because the best models changed little when fire years were included, it can be inferred from the regression analyses that fire, as applied during the month of May, did not greatly affect how the number of flowering orchids responded to precipitation in the life stages evaluated. The ANCOVA analyses provided complementary results, finding no significant effect of fire on the number of flowering orchids after taking into account the variability in precipitation.

Shorter plants were observed in years in which prescribed fire occurred, although most of the burns occurred in years with low spring precipitation. Pleasants (1994) conducted a controlled burning experiment and hypothesized that growth in *P. praeclara* may be affected by an interaction of fire and soil moisture, with fire in dry years suppressing growth rates, and fire in wet years increasing growth rates. His study was limited, however, in that it only included burns in two years. While our results are supportive of this hypothesis, a more parsimonious interpretation of our data is that precipitation was important in determining plant height, but fire was not, as the effect of fire disappeared when variability in precipitation was taken into account. We also found no evidence for a one year lag in the effect of fire.

Thus, overall, our data support important ecological effects of precipitation but no effects of fire applied in May on orchid growth or reproduction. An important caveat is that if fire intensity, fire timing,

and precipitation do interact (and these factors may exhibit considerable variability), then a much larger sample size would be required to evaluate such interactions.

Land managers face a dilemma when using prescribed fire to maintain habitat occupied by *P. praeclara*, while preventing direct injury to orchids. The prescribed fires during May evaluated in this study were timed to reduce the abundance of the exotic cool season grass *Bromus inermis* Leyss (smooth brome; Willson and Stubbendieck 1997). Late spring fires may damage *P. praeclara* individuals likely to flower, however, and fall burns provide an alternative strategy for reducing smooth brome (Biederman et al. 2014). While no effect of fire was observed on the orchid populations of our study, relatively few burns occurred over the course of the data record, and no control plots existed. The timing of emergence of *P. praeclara* is variable, and spring rains often make burning difficult. Thus, burns in October and November, when the orchid is dormant, may provide better opportunities to introduce fire without impact to *P. praeclara*, although this may increase the probability of smooth brome expansion.

ACKNOWLEDGMENTS

Views, statements, findings, conclusions, recommendations, and data in this report are those of the author(s) and do not necessarily reflect views and policies of the National Park Service, US Department of the Interior. Mention of trade names or commercial products does not constitute endorsement or recommendation for use by the National Park Service. We thank G.D. Willson and F.A. Akyüz for early development work on the protocol and data collection, as well as Pipestone National Monument staff who have assisted over the years, especially K. Legg, G. Wagner, B. Hummel, and S. Hendriks. S. Leis provided helpful comments on the manuscript.

All co-authors are ecologists with the National Park Service's Heartland Network Inventory and Monitoring Program.

Lloyd Morrison, quantitative ecologist, has

research interests in island biogeography, metapopulation dynamics and invasion biology.

Jennifer Haack-Gaynor, GIS specialist, focuses on research in land cover and prairie vegetation monitoring.

Craig Young specializes in invasive plant monitoring and management.

Mike DeBacker, network coordinator, has focused on the disturbance ecology of prairies in the Great Plains.

LITERATURE CITED

- Akaike, H. 1973. Information theory as an extension of the maximum likelihood principle. Pp. 267-281 in B.N. Petrov and F. Csaki, eds., Second International Symposium on Information Theory. Akademiai Kiado, Budapest.
- Anderson, D.R., and K.P. Burnham. 2002. Avoiding pitfalls when using information – theoretic methods. *Journal of Wildlife Management* 66:912-918.
- Austin, P.C., and J.V. Tu. 2004. Automated variable selection methods for logistic regression produced unstable models for predicting acute myocardial infarction mortality. *Journal of Clinical Epidemiology* 57:1138-1146.
- Becker, D.A., T.B. Bragg, and D.M. Sutherland. 1986. Vegetation survey and prairie management plan for Pipestone National Monument. National Park Service, Midwest Regional Office, Omaha, NE.
- Biederman, L.A., J. Beckman, J. Prekker, D. Anderson, N.P. Sather, and R. Dahle. 2014. Phenological monitoring aids habitat management of threatened plant. *Natural Areas Journal* 34:105-110.
- Bowles, M.L. 1983. The tallgrass prairie orchids *Platanthera leucophaea* (Nutt.) Lindl. and *Cypripedium candidum* Muhl. ex Willd.: Some aspects of their status, biology, and ecology, and implications toward management. *Natural Areas Journal* 3:14-37.
- Bowles, M.L., R. Flakne, and R. Dombeck. 1992. Status and population fluctuations of the eastern prairie fringed orchid [*Platanthera leucophaea* (Nutt.) Lindl.] in Illinois. *Erigenia* 12:26-40.
- Bray, T.E., and B.L. Wilson. 1992. Status of *Platanthera praeclara* Sheviak and Bowles (western prairie fringed orchid) in the Platte River Valley in Nebraska from Hamilton to

- Garden counties. Transactions of the Nebraska Academy of Sciences 19:57-62.
- Burnham, K.P., and D.R. Anderson. 2002. Model Selection and Multi-model Inference. 2nd ed. Springer-Verlag, New York.
- Currier, P.J. 1984. Response of prairie fringed orchid to fire and reduction in grazing (Nebraska). Restoration and Management Notes 2:28.
- Hegyi, G., and L.Z. Garamszegi. 2011. Using information theory as a substitute for stepwise regression in ecology and behavior. Behavioral Ecology and Sociobiology 65:69-76.
- IBM Corp. 2011. IBM SPSS Statistics for Windows, Version 20.0. IBM Corp., Armonk, NY.
- Kullback, S., and R.A. Leibler. 1951. On information and sufficiency. Annals of Mathematical Statistics 22:79-86.
- Kutner, M.H., C.J. Nachtsheim, J. Neter, and W. Li. 2005. Applied Linear Statistical Models. McGraw Hill Irwin, Boston, MA.
- Pleasants, J.M. 1994. The effects of spring burns on the western prairie fringed orchid (*Platanthera praeclara*). Pp. 67-73 in D.C. Hartnett, ed., Proceedings of the 14th Annual North American Prairie Conference. Kansas State University, Manhattan.
- Rice, W.R. 1989. Analyzing tables of statistical tests. Evolution 43:223-225.
- SAS Institute Inc. 2008. SAS for Windows. Version 9.2. SAS Institute Inc. Cary, NC.
- Sather, N., and D. Anderson. 2010. Twenty-five years of monitoring the western prairie fringed orchid (*Platanthera* Sheviak and Bowles) in Minnesota. Pp. 126-134 in D. Williams, B. Butler, and D. Smith, eds., 22nd North American Prairie Conference. Tallgrass Prairie Center, University of Northern Iowa, Cedar Falls.
- Sharma, J., L.W. Zettler, J.W. Van Sambeek, M.R. Ellersieck, and C.J. Starbuck. 2003. Symbiotic seed germination and mycorrhizae of federally threatened *Platanthera praeclara* (Orchidaceae). American Midland Naturalist 149:104-120.
- Sheviak, C.J. 1974. An introduction to the ecology of the Illinois Orchidaceae. Illinois State Museum Scientific Papers 14:1-89.
- Sheviak, C.J., and M.L. Bowles. 1986. The prairie fringed orchids: a pollinator-isolated species pair. Rhodora 88:267-290.
- Sieg, C.H., and R.M. King. 1995. Influence of environmental factors and preliminary demographic analyses of a threatened orchid, *Platanthera leucophaea*. American Midland Naturalist 134:307-323.
- US Fish and Wildlife Service. 1989. Endangered and threatened wildlife and plants; determination of threatened status for *Platanthera leucophaea* (eastern prairie fringed orchid) and *Platanthera praeclara* (western prairie fringed orchid). Federal Register 54:39857-39862.
- US Fish and Wildlife Service. 1996. *Platanthera praeclara* (western prairie fringed orchid) recovery plan. US Fish and Wildlife Service, Fort Snelling, MN.
- Whittingham, M.J., P.A. Stephens, R.B. Bradbury, and R.P. Freckleton. 2006. Why do we still use stepwise modelling in ecology and behaviour? Journal of Animal Ecology 75:1182-1189.
- Wiegand, R.E. 2010. Performance of using multiple stepwise algorithms for variable selection. Statistics in Medicine 29:1647-1659.
- Willson, G., M. Page, and F.A. Akyuz. 2006. Precipitation and fire effects on flowering of a rare prairie orchid. Great Plains Research 16:37-43.
- Willson, G.D. 2000. Western prairie fringed orchid monitoring protocol for Pipestone National Monument. US Geological Survey, US Department of the Interior, Northern Prairie Wildlife Research Center, Columbia, MO.
- Willson, G.D., and J. Stubbendieck. 1997. Fire effects on four growth stages of smooth brome (*Bromus inermis* Leyss.). Natural Areas Journal 17:306-312.
- Wolken, P.M. 1995. Habitat and life history of the western prairie fringed orchid (*Platanthera praeclara*). M.S. thesis, University of Wyoming, Laramie.
- Young, C.C., G.D. Willson, M.D. DeBacker, L.W. Morrison, H.J. Etheridge, and J.L. Haack. 2007. Western Prairie Fringed Orchid Monitoring Protocol for Pipestone National Monument. Natural Resource Report NPS/HTLN/NRR-2007/013. National Park Service, Fort Collins, CO.

Appendix. Data set, including precipitation during various phenological stages, used in analyses.

Year	No. of orchids	No. of flowers (mean \pm SE)	Plant height (cm) (mean \pm SE)	Precipitation (cm) during phenological stage ^a :					
				MGprev	SB	PS	DO	EM	MGcurr
1993	33	9.4 \pm 0.6	59.6 \pm 1.4	21.79	14.17	7.06	18.90	27.43	41.02
1994 ^b	18	8.8 \pm 1.1	57.3 \pm 2.5	41.02	7.42	7.26	13.92	15.95	14.25
1995	37	9.5 \pm 0.5	61.4 \pm 1.2	14.25	9.40	8.18	22.17	24.28	23.5
1996	55	8.9 \pm 0.5	60.0 \pm 1.2	23.50	8.15	8.89	18.34	11.28	11.13
1997 ^b	3	9.3 \pm 1.9	38.7 \pm 3.8	11.13	5.54	11.63	27.25	10.39	14.17
1998	0			14.17	2.82	4.88	20.83	10.67	14.30
1999	16	7.9 \pm 0.7	55.1 \pm 2.6	14.30	8.33	3.51	31.75	21.51	25.83
2000	125	8.1 \pm 0.3	47.2 \pm 0.7	25.83	5.82	3.43	9.32	27.53	9.63
2001	95	7.7 \pm 0.2	51.3 \pm 0.8	9.63	6.07	2.41	25.10	26.57	19.76
2002 ^b	124	8.4 \pm 0.3	37.5 \pm 0.6	19.76	2.03	7.98	18.65	9.98	17.22
2003	221	8.3 \pm 0.2	55.9 \pm 0.6	17.22	19.46	5.00	14.05	26.37	14.91
2004	146	7.3 \pm 0.2	53.8 \pm 0.7	14.91	4.24	10.90	15.98	19.89	20.24
2005	149	7.0 \pm 0.2	52.8 \pm 0.8	20.24	7.24	13.23	8.76	15.16	29.49
2006	100	7.6 \pm 0.2	44.7 \pm 0.7	29.49	5.99	20.07	23.27	12.37	7.77
2007	1	5.0	40.0	7.77	13.89	8.74	16.21	13.97	3.45
2008	10	7.2 \pm 0.5	49.8 \pm 2.0	3.45	15.75	6.55	21.36	14.33	19.51
2009 ^b	28	6.1 \pm 0.4	38.2 \pm 1.1	19.51	4.60	3.84	22.25	9.47	17.22
2010	92	7.6 \pm 0.2	52.4 \pm 0.8	17.22	10.16	2.26	26.37	14.07	35.05
2011	645	9.1 \pm 0.2	55.1 \pm 0.4	29.87	12.85	24.33	18.24	24.74	29.87
2012	722	9.0 \pm 0.1	51.6 \pm 0.3	35.05	3.78	0.58	12.12	40.56	5.33

^a MGprev = mature growth in the year before flowering (June 1 – July 31 of previous year); SB = senescence/bud development (August 1 – 31 of previous year); PS = post-senescence (September 1 – 30 of previous year); DO = dormancy (October 1 – December 31 of previous year and January 1 – March 31 of current year); EM = emergence (April 1 – May 31 of current year); MGcurr = mature growth in the year of flowering (June 1 – July 31 of current year).

^b Denotes year in which a prescribed fire occurred.